# A Tracy-Widom Empirical Estimator For Valid P-values With High-Dimensional Datasets 

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## Motivating Example

## Systemic Autoimmune Diseases

- Systemic Autoimmune diseases, e.g. Rheumatoid arthritis, Lupus, Scleroderma, impact many systems at once.
- We want to study the association between DNA methylation and these diseases
- To account for the complex biological architecture, we want to measure the association at the genetic pathway level
- High-Dimensional Data

How can we efficiently compute valid p-values?

High-dimensional inference

## Double Wishart Problem

- Many multivariate methods involve maximising a Rayleigh quotient:

$$
R^{2}(w)=\frac{w^{\top} A w}{w^{\top}(A+B) w}
$$

- This approach is equivalent to finding the largest root $\lambda$ of a double Wishart problem:

$$
\operatorname{det}(\mathbf{A}-\lambda(\mathbf{A}+\mathbf{B}))=0
$$

## Double Wishart Problem

Well-known examples of double Wishart problems:

- Multivariate Analysis of Variance (MANOVA);
- Canonical Correlation Analysis (CCA);
- Testing for independence of two multivariate samples;
- Testing for the equality of covariance matrices of two independent samples from multivariate normal distributions;

In all the examples above, the largest root $\lambda$ summarises the strength of the association.

## Contributions

The main contribution:

1. I will provide an empirical estimate of the distribution of the largest root of the determinantal equation. This estimate can be used to compute valid $p$-values and perform high-dimensional inference.

Two $R$ packages implement this method: pcev and covequal (both available on CRAN)

## Inference

There is evidence in the literature that the null distribution of the largest root $\lambda$ should be related to the Tracy-Widom distribution.

Theorem
(Johnstone 2008) Assume $\mathbf{A} \sim W_{p}(\Sigma, m)$ and $\mathbf{B} \sim W_{p}(\Sigma, n)$ are independent, with $\Sigma$ positive-definite and $\boldsymbol{n} \leq \boldsymbol{p}$. As $p, m, n \rightarrow \infty$, we have

$$
\frac{\operatorname{logit} \lambda-\mu}{\sigma} \xrightarrow{\mathcal{D}} T W(1),
$$

where $T W(1)$ is the Tracy-Widom distribution of order 1 , and $\mu, \sigma$ are explicit functions of $p, m, n$.

## Inference

- However, Johnstone's theorem requires an invertible matrix.
- The null distribution of $\lambda$ is asymptotically equal to that of the largest root of a scaled Wishart (Srivastava).
- The null distribution of the largest root of a Wishart is also related to the Tracy-Widom distribution.
- More generally, random matrix theory suggests that the Tracy-widom distribution is key in central-limit-like theorems for random matrices.


## Empirical Estimate

We propose to obtain an empirical estimate as follows:

## Estimate the null distribution

1. Perform a small number of permutations ( $\sim 50$ ).

- The actual procedure is problem-specific.

2. For each permutation, compute the largest root statistic.
3. Fit a location-scale variant of the Tracy-Widom distribution.

Numerical investigations support this approach for computing $\mathbf{p}$-values. The main advantage over a traditional permutation strategy is the computation time.

## Simulations

## Distribution Estimation

- We generated 1000 pairs of Wishart variates $\mathbf{A} \sim W_{p}(\Sigma, m)$, $\mathbf{B} \sim W_{p}(\Sigma, n)$ with $m=96$ and $n=4$ fixed
- MANOVA: this would correspond to four distinct populations and a total sample size of 100
- We varied $p=500,1000,1500,2000$
- We looked at two different covariance structures: $\Sigma=I_{p}$, and an exchangeable correlation structure with parameter $\rho=0.2$.
- We looked at four different numbers of permutations for the empirical estimator: $K=25,50,75,100$.
- We compared graphically the CDF estimated from the empirical estimate with the true CDF


## Distribution Estimation

$$
\text { Type - True CDF - Heuris. } 25 \text { - Heuris. } 50 \text { - Heuris. } 75 \text { - Heuris. } 100
$$



## P-value Comparison

We looked at the following high-dimensional simulation scenario:

- We fixed $n=100$.
- We generated $X \sim N_{p}\left(0, I_{p}\right)$ and $\mathbf{Y} \sim N_{p}(0, \Sigma)$, with $p=200,300,400,500$.
- We selected an autocorrelation structure $\Sigma$ :

$$
\operatorname{Cov}\left(Y_{i}, Y_{j}\right)=\rho^{|i-j|}, \quad \rho=0,0.2
$$

- We compared the empirical estimate with a permutation procedure (250 permutations).
- Each simulation was repeated 100 times.


## P-value Comparison



## Data Analysis

- DNA methylation measured with Illumina 450k on 28 cell-separated samples
- We focus on Monocytes only.
- 18 patients suffering from Rheumatoid arthritis, Lupus, Scleroderma
- We group locations by biological KEGG pathways
- The number of genomic locations per pathway ranged from 39 to 21,640 , with an average around 2000 dinucleotides.
- 134,941 CpG dinucleotides were successfully matched to one of 320 KEGG pathways
- On average, each locations appears in 4.5 pathways $\Rightarrow$ effectively 70 independent hypothesis tests


## Results

| Description | P-value | P-value (permutation) |
| :--- | :---: | :---: |
| Glutamatergic synapse | $1.91 \times 10^{-4}$ | $7.00 \times 10^{-4}$ |
| Ras signaling pathway | $1.33 \times 10^{-3}$ | $1.40 \times 10^{-3}$ |
| Circadian rhythm | $1.52 \times 10^{-3}$ | $1.00 \times 10^{-4}$ |
| Histidine metabolism | $1.59 \times 10^{-3}$ | $3.00 \times 10^{-4}$ |
| Pathogenic E. coli infection | $1.65 \times 10^{-3}$ | $5.20 \times 10^{-3}$ |

## Results


path:hsa00120—Glutamatergic synapse: Comparison of VIF and univariate p -values for the most significant pathway.

## Conclusion

- Data summary is an important feature in data analysis, and this is the objective of dimension reduction techniques.
- In a high-dimensional setting, estimation and inference are more challenging
- Estimation: Truncated SVD
- Inference: Fitted location-scale Tracy-Widom
- Our approach is computationally simple.
- Everything presented today has been implemented in two $R$ packages.

Demo

## Principal Component of Explained Variance (PCEV)

- Provides an optimal strategy for selecting a low dimensional summary of $\mathbf{Y}$ that can be used to test for association with one or several covariates of interest.
- Goal: Find the linear combination (or component) that maximises the proportion of variance explained by the covariates


## PCEV: Statistical Model

Let $\mathbf{Y}$ be a multivariate outcome of dimension $p$ and $X$, a vector of covariates.

We assume a linear relationship:

$$
\mathbf{Y}=\beta^{T} X+\varepsilon
$$

The total variance of the outcome can then be decomposed as

$$
\begin{aligned}
\operatorname{Var}(\mathbf{Y}) & =\operatorname{Var}\left(\beta^{T} X\right)+\operatorname{Var}(\varepsilon) \\
& =V_{M}+V_{R}
\end{aligned}
$$

## PCEV: Statistical Model

Decompose the total variance of $\mathbf{Y}$ into:

1. Variance explained by the covariates;
2. Residual variance.

## PCEV: Statistical Model

The PCEV framework seeks a linear combination $w^{\top} \mathbf{Y}$ such that the proportion of variance explained by $X$ is maximised; this proportion is defined as the following Rayleigh quotient:

$$
R^{2}(w)=\frac{w^{\top} V_{M} w}{w^{T}\left(V_{M}+V_{R}\right) w} .
$$

A solution to this maximisation problem can be obtained through a combination of Lagrange multipliers and linear algebra.

Key observation: $R^{2}(w)$ measures the strength of the association

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## Questions or comments?

## For more information and updates, visit <br> maxturgeon.ca.

