# A Tracy-Widom Empirical Estimator For Valid P-values With High-Dimensional Datasets

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## **Motivating Example**

#### Systemic Autoimmune Diseases

- Systemic Autoimmune diseases, e.g. Rheumatoid arthritis, Lupus, Scleroderma, impact many systems at once.
- We want to study the association between DNA methylation and these diseases
- To account for the complex biological architecture, we want to measure the association at the *genetic pathway level*
- High-Dimensional Data

How can we efficiently compute valid p-values?

## **High-dimensional inference**

• Many multivariate methods involve maximising a Rayleigh quotient:

$$R^2(w) = \frac{w^T A w}{w^T (A+B) w}$$

 This approach is equivalent to finding the largest root λ of a double Wishart problem:

$$\det \left( \mathbf{A} - \lambda (\mathbf{A} + \mathbf{B}) \right) = 0.$$

Well-known examples of double Wishart problems:

- Multivariate Analysis of Variance (MANOVA);
- Canonical Correlation Analysis (CCA);
- Testing for independence of two multivariate samples;
- Testing for the equality of covariance matrices of two independent samples from multivariate normal distributions;

In all the examples above, the largest root  $\lambda$  summarises the strength of the association.

The main contribution:

 I will provide an empirical estimate of the distribution of the largest root of the determinantal equation. This estimate can be used to compute valid p-values and perform high-dimensional inference.

Two R packages implement this method: pcev and covequal (both available on CRAN)

There is evidence in the literature that the null distribution of the largest root  $\lambda$  should be related to the **Tracy-Widom distribution**.

#### Theorem

(Johnstone 2008) Assume  $\mathbf{A} \sim W_p(\Sigma, m)$  and  $\mathbf{B} \sim W_p(\Sigma, n)$  are independent, with  $\Sigma$  positive-definite and  $\mathbf{n} \leq \mathbf{p}$ . As  $p, m, n \to \infty$ , we have

$$\frac{\operatorname{logit} \lambda - \mu}{\sigma} \xrightarrow{\mathcal{D}} TW(1),$$

where TW(1) is the Tracy-Widom distribution of order 1, and  $\mu, \sigma$  are explicit functions of p, m, n.

- However, Johnstone's theorem requires an invertible matrix.
- The null distribution of  $\lambda$  is asymptotically equal to that of the largest root of a scaled Wishart (Srivastava).
  - The null distribution of the largest root of a Wishart is also related to the Tracy-Widom distribution.
- More generally, random matrix theory suggests that the Tracy-widom distribution is key in central-limit-like theorems for random matrices.

We propose to obtain an empirical estimate as follows:

#### Estimate the null distribution

- 1. Perform a small number of permutations ( $\sim$  50).
  - The actual procedure is problem-specific.
- 2. For each permutation, compute the largest root statistic.
- 3. Fit a location-scale variant of the Tracy-Widom distribution.

Numerical investigations support this approach for computing p-values. The main advantage over a traditional permutation strategy is the computation time.

## Simulations

#### **Distribution Estimation**

- We generated 1000 pairs of Wishart variates A ~ W<sub>p</sub>(Σ, m), B ~ W<sub>p</sub>(Σ, n) with m = 96 and n = 4 fixed
  - MANOVA: this would correspond to four distinct populations and a total sample size of 100
- We varied *p* = 500, 1000, 1500, 2000
- We looked at two different covariance structures: Σ = I<sub>p</sub>, and an exchangeable correlation structure with parameter ρ = 0.2.
- We looked at four different numbers of permutations for the empirical estimator: K = 25, 50, 75, 100.
- We compared graphically the CDF estimated from the empirical estimate with the true CDF

#### **Distribution Estimation**



We looked at the following high-dimensional simulation scenario:

- We fixed n = 100.
- We generated  $X \sim N_{p}(0, I_{p})$  and  $\mathbf{Y} \sim N_{p}(0, \Sigma)$ , with p = 200, 300, 400, 500.
- We selected an autocorrelation structure  $\Sigma$ :

$$Cov(Y_i, Y_j) = \rho^{|i-j|}, \qquad \rho = 0, 0.2$$

- We compared the empirical estimate with a permutation procedure (250 permutations).
- Each simulation was repeated 100 times.

#### **P-value Comparison**



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Data Analysis

- DNA methylation measured with Illumina 450k on 28 cell-separated samples
- We focus on Monocytes only.
- 18 patients suffering from Rheumatoid arthritis, Lupus, Scleroderma
- We group locations by biological KEGG pathways
  - The number of genomic locations per pathway ranged from 39 to 21,640, with an average around 2000 dinucleotides.
  - 134,941 CpG dinucleotides were successfully matched to one of 320 KEGG pathways
  - On average, each locations appears in 4.5 pathways  $\Rightarrow$  effectively 70 independent hypothesis tests

Description	P-value	P-value (permutation)
Glutamatergic synapse	$1.91  imes 10^{-4}$	$7.00 imes10^{-4}$
Ras signaling pathway	$1.33\times10^{-3}$	$1.40 imes10^{-3}$
Circadian rhythm	$1.52  imes 10^{-3}$	$1.00 imes10^{-4}$
Histidine metabolism	$1.59 imes10^{-3}$	$3.00 imes10^{-4}$
Pathogenic E. coli infection	$1.65 imes10^{-3}$	$5.20 imes10^{-3}$

#### Results



**path:hsa00120—Glutamatergic synapse**: Comparison of VIF and univariate p-values for the most significant pathway.

- Data summary is an important feature in data analysis, and this is the objective of dimension reduction techniques.
- In a high-dimensional setting, **estimation** and **inference** are more challenging
  - Estimation: Truncated SVD
  - Inference: Fitted location-scale Tracy-Widom
- Our approach is computationally simple.
- Everything presented today has been implemented in two R packages.

## Demo

- Provides an optimal strategy for selecting a low dimensional summary of Y that can be used to test for association with one or several covariates of interest.
- **Goal**: Find the linear combination (or component) that maximises the *proportion of variance explained by the covariates*

#### **PCEV: Statistical Model**

Let **Y** be a multivariate outcome of dimension p and X, a vector of covariates.

We assume a linear relationship:

$$\mathbf{Y} = \beta^T X + \varepsilon.$$

The total variance of the outcome can then be decomposed as

$$Var(\mathbf{Y}) = Var(\beta^T X) + Var(\varepsilon)$$
$$= V_M + V_R.$$

## Decompose the total variance of $\boldsymbol{\mathsf{Y}}$ into:

- 1. Variance explained by the covariates;
- 2. Residual variance.

The PCEV framework seeks a linear combination  $w^T \mathbf{Y}$  such that the proportion of variance explained by X is maximised; this proportion is defined as the following Rayleigh quotient:

$$R^2(w) = \frac{w^T V_M w}{w^T (V_M + V_R) w}.$$

A solution to this maximisation problem can be obtained through a combination of Lagrange multipliers and linear algebra.

**Key observation**:  $R^2(w)$  measures the strength of the association

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#### **Questions or comments?**

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